

Causal Inference for Survival Outcomes: An Introduction

Bianca L De Stavola for STRATOS TG 7

Great Ormond Street Institute of Child Health, University College London

LSHTM Centre for Statistical Methodology , 3rd Nov 2021



Topic group 7: Causal Inference

Saskia le Cessie
Bianca De Stavola
Vanessa Didelez
Els Goetghebeur
Erica Moodie
Ingeborg Waernbaum

Leiden University Medical Centre & Ghent University
University College London, UK
University of Bremen, Germany
Ghent University, Belgium
McGill University, Canada
Uppsala University, Sweden

ofcaus.org





- ▶ To find a causal answer, start with a causal question.

Then:

- 1 specify exposure, outcome, population of interest, target causal effects (**estimands**)
- 2 state assumptions for identification and estimation of effects from the data
- 3 interpret results cautiously.

[Goetghebeur *et al.* 2020]



- ▷ To find a causal answer, start with a causal question.

Then:

- 1 specify exposure, outcome, population of interest, target causal effects (**estimands**)
- 2 state assumptions for identification and estimation of effects from the data
- 3 interpret results cautiously.

[Goetghebeur *et al.* 2020]

- ▷ These are indeed the principles guiding RCTs and, for observational studies, are referred to as "**target trial emulation**" (TTE).

Estimands

- ▶ Accepted specification of estimands is via **potential outcomes**
- ▶ Y_a : outcome that would arise had we set A to take the value a by a well-defined (hypothetical) intervention.

Estimands

- ▶ Accepted specification of estimands is via **potential outcomes**
- ▶ Y_a : outcome that would arise had we set A to take the value a by a well-defined (hypothetical) intervention.
- ▶ For example (for a binary exposure A):
 - $ACE = E(Y_1) - E(Y_0)$
 - $ATT = E(Y_1|A = 1) - E(Y_0|A = 1)$

Estimands

- ▶ Accepted specification of estimands is via **potential outcomes**
- ▶ Y_a : outcome that would arise had we set A to take the value a by a well-defined (hypothetical) intervention.
- ▶ For example (for a binary exposure A):
 - $ACE = E(Y_1) - E(Y_0)$
 - $ATT = E(Y_1|A = 1) - E(Y_0|A = 1)$
- ▶ Identification requires **linking data to these hypothetical quantities**, e.g. invoking assumptions of no interference, consistency, and positivity.

Estimands

- ▶ Accepted specification of estimands is via **potential outcomes**
- ▶ Y_a : outcome that would arise had we set A to take the value a by a well-defined (hypothetical) intervention.
- ▶ For example (for a binary exposure A):
 - $ACE = E(Y_1) - E(Y_0)$
 - $ATT = E(Y_1|A = 1) - E(Y_0|A = 1)$
- ▶ Identification requires **linking data to** these **hypothetical quantities**, e.g. invoking assumptions of no interference, consistency, and positivity.
- ▶ Choice of **estimation methods**: each requiring additional assumptions (e.g. no unmeasured confounding, correct (semi-)parametric models).

Estimands for time-to-event outcomes

- ▷ Well known challenges [Andersen *et al.* 2020]:
 - **Time origin** and scale: from birth/entry/surgery?
 - **Censoring**: information on whether event is observed.

Estimands for time-to-event outcomes

- ▷ Well known challenges [Andersen *et al.* 2020]:
 - **Time origin** and scale: from birth/entry/surgery?
 - **Censoring**: information on whether event is observed.

Which aspects of time should we focus on when comparing alternative worlds?

Outline

- 1 Basic Principles
- 2 Causality and Survival Analysis
 - Estimands
 - Dealing with censoring
- 3 Identification and Estimation
 - Dealing with censoring
- 4 The Simulation Learner
 - Results
- 5 Summary
 - References

Survival Estimands (Version I)

Let T_a be the **potential survival time** if we set A to take the value a by a well-defined (hypothetical) intervention.

Survival Estimands (Version I)

Let T_a be the **potential survival time** if we set A to take the value a by a well-defined (hypothetical) intervention.

(A) **Risk scale:**

Differences in survival probabilities at relevant times

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

This is the difference in (marginal) survival functions of POs,

Survival Estimands (Version I)

Let T_a be the **potential survival time** if we set A to take the value a by a well-defined (hypothetical) intervention.

(A) **Risk scale:**

Differences in survival probabilities at relevant times

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

This is the difference in (marginal) survival functions of POs,

▷ **Interpretation:**

risk difference for no event by time t had random patient been treated versus not.

Survival Estimands (Version I)

(B) Hazard scale:

Contrast of hazards, for example $\frac{\lambda_1(t)}{\lambda_0(t)}$, where

$$\lambda_a(t) = \lim_{h \rightarrow 0} \frac{1}{h} P(t \leq T_a < t + h | T_a \geq t)$$

Survival Estimands (Version I)

(B) Hazard scale:

Contrast of hazards, for example $\frac{\lambda_1(t)}{\lambda_0(t)}$, where

$$\lambda_a(t) = \lim_{h \rightarrow 0} \frac{1}{h} P(t \leq T_a < t + h | T_a \geq t)$$

▷ **Complication:**

Interpretational difficulties because of the built-in selection due to the conditioning on different subgroups ($T_0 \geq t$ and $T_1 \geq t$).

Survival Estimands (Version I)

(C) Restricted Mean Survival Times (RMST):

$$\int_0^{\tau} P(T_1 > t) dt - \int_0^{\tau} P(T_0 > t) dt$$

These mean times are specific to τ

Survival Estimands (Version I)

(C) Restricted Mean Survival Times (RMST):

$$\int_0^{\tau} P(T_1 > t)dt - \int_0^{\tau} P(T_0 > t)dt$$

These mean times are specific to τ

(D) Mean Survival Time:

$$\int_0^{\infty} P(T_1 > t)dt - \int_0^{\infty} P(T_0 > t)dt$$

Preferred in econometrics, e.g. in `stteffects` in Stata

Survival Estimands (Version I)

(C) **Restricted Mean Survival Times (RMST):**

$$\int_0^{\tau} P(T_1 > t) dt - \int_0^{\tau} P(T_0 > t) dt$$

These mean times are specific to τ

(D) **Mean Survival Time:**

$$\int_0^{\infty} P(T_1 > t) dt - \int_0^{\infty} P(T_0 > t) dt$$

Preferred in econometrics, e.g. in `stteffects` in Stata

(E) **Other scales**, e.g. speed from Accelerated Failure Time Models.

Comments

- ▷ The choice between these estimands should be guided by their clinical relevance.
- ▷ In most settings these are contrasts on risk scale.
- ▷ Note however that hazard models are useful to derive such contrasts.

Comments

- ▷ The choice between these estimands should be guided by their clinical relevance.
- ▷ In most settings these are contrasts on risk scale.
- ▷ Note however that hazard models are useful to derive such contrasts.
- ▷ Whichever one is chosen, definitions above have no consideration of the impact of censoring.

Censoring

- Do we want to quantify causal effects in the absence in censoring?

Censoring

- Do we want to quantify causal effects in the absence in censoring?

It depends on the source of censoring: for some it does not always make clinical sense to remove them,

- Administrative reasons
- Loss to follow-up
- Treatment switching
- Competing event

Censoring

- Do we want to quantify causal effects in the absence in censoring?

It depends on the source of censoring: for some it does not always make clinical sense to remove them,

- Administrative reasons ✓
- Loss to follow-up ✓
- Treatment switching
- Competing event

Censoring

- Do we want to quantify causal effects in the absence in censoring?

It depends on the source of censoring: for some it does not always make clinical sense to remove them,

- Administrative reasons
- Loss to follow-up
- Treatment switching ??
- Competing event ??

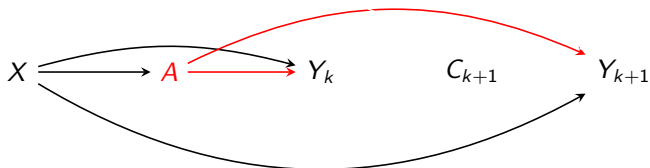
An Introduction

└ Causality and Survival Analysis

└ Dealing with censoring

The problem with censoring

Consider a subset of the follow-up with **administrative** censoring:



[Adapted from Young *et al.*, 2020]

A: baseline treatment

Y_k : outcome at time t_k

X : baseline confounders

C_k : censoring indicator at time t_k

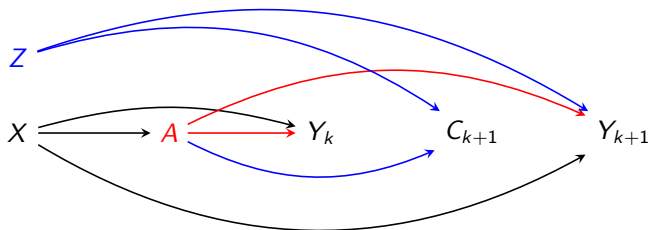
An Introduction

└ Causality and Survival Analysis

└ Dealing with censoring

The problem with censoring

Consider a subset of the follow-up with **informative** censoring:



[Adapted from Young *et al.*, 2020]

A: baseline treatment

Y_k : outcome at time t_k

X : baseline confounders

C_k : censoring indicator at time t_k **Z**: predictors of censoring and outcome

Survival Estimands (Version II)

- ▷ If happy to phrase the question “in the absence of censoring”,
- ▷ Risk scale:

$$\text{ACE}(t) = P(T_{A=1, C=0} > t) - P(T_{A=0, C=0} > t), \quad t \text{ in } [0, \tau]$$

where $T_{A=a, C=0}$ is the potential survival time if we set A to take the value a and $C = 0$ over the entire follow-up.

Survival Estimands (Version II)

- ▷ If happy to phrase the question “in the absence of censoring”,
- ▷ Risk scale:

$$\text{ACE}(t) = P(T_{A=1, C=0} > t) - P(T_{A=0, C=0} > t), \quad t \text{ in } [0, \tau]$$

where $T_{A=a, C=0}$ is the potential survival time if we set A to take the value a and $C = 0$ over the entire follow-up.

- ▷ and similarly for the other estimands.

Identification

For identification of causal effects for a time-to-event outcome, invoke:

- ▶ No interference, consistency and positivity of the exposure
- ▶ No unmeasured confounding (NUC), *i.e.* : sufficient covariate information regarding treatment assignment confounding

Identification

For identification of causal effects for a time-to-event outcome, invoke:

- ▶ No interference, consistency and positivity of the exposure
- ▶ No unmeasured confounding (NUC), *i.e.* : sufficient covariate information regarding treatment assignment confounding
- ▶ In the presence of censoring we also require: sufficient covariate information regarding (possibly time-varying) 'common causes' of censoring and event.

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

1. Model-based marginal counterfactual survival curves:

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

1. Model-based marginal counterfactual survival curves:

- (Sufficiently) flexible survival models (e.g. proportional or additive hazard model)

Estimation of ACE(t)

With administrative censoring

$$\text{ACE}(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

1. Model-based marginal counterfactual survival curves:

- (Sufficiently) flexible survival models (e.g. proportional or additive hazard model)
- Derive individual-level predicted potential survival curves

Estimation of ACE(t)

With administrative censoring

$$\text{ACE}(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

1. Model-based marginal counterfactual survival curves:

- (Sufficiently) flexible survival models (e.g. proportional or additive hazard model)
- Derive individual-level predicted potential survival curves
- Standardisation to the distribution of the observed confounders (*“empirical standardization”*)

Estimation of ACE(t)

With administrative censoring

$$\text{ACE}(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

1. Model-based marginal counterfactual survival curves:

- (Sufficiently) flexible survival models (e.g. proportional or additive hazard model)
- Derive individual-level predicted potential survival curves
- Standardisation to the distribution of the observed confounders (*“empirical standardization”*)
- Compute difference at selected values of t .

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

- Fit propensity score (PS) model for treatment A

Estimation of ACE(t)

With administrative censoring

$$\text{ACE}(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

- Fit propensity score (PS) model for treatment A
- Derive stabilized weights w from the predicted PS (IPTW)

Estimation of ACE(t)

With administrative censoring

$$\text{ACE}(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

- Fit propensity score (PS) model for treatment A
- Derive stabilized weights w from the predicted PS (IPTW)
- Fit survival MSM for treatment by reweighing the data using w

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

- Fit propensity score (PS) model for treatment A
- Derive stabilized weights w from the predicted PS (IPTW)
- Fit survival MSM for treatment by reweighing the data using w
- Use the model to predict potential survival curves

Estimation of $ACE(t)$

With administrative censoring

$$ACE(t) = P(T_1 > t) - P(T_0 > t), \quad t \text{ in } [0, \tau]$$

2. IPTW of Marginal Structural Models (MSM):

- Fit propensity score (PS) model for treatment A
- Derive stabilized weights w from the predicted PS (IPTW)
- Fit survival MSM for treatment by reweighing the data using w
- Use the model to predict potential survival curves
- Compute difference at selected values of t .

Estimation of $ACE(t)$

With informative censoring

$$ACE(t) = P(T_{A=1, C=0} > t) - P(T_{A=0, C=0} > t), \quad t \text{ in } [0, \tau]$$

Estimation requires dealing with the drivers of censoring:

Estimation of $ACE(t)$

With informative censoring

$$ACE(t) = P(T_{A=1, C=0} > t) - P(T_{A=0, C=0} > t), \quad t \text{ in } [0, \tau]$$

Estimation requires dealing with the drivers of censoring:

Approach 1. Include predictors of missingness among the conditioning factors

Estimation of $ACE(t)$

With informative censoring

$$ACE(t) = P(T_{A=1, C=0} > t) - P(T_{A=0, C=0} > t), \quad t \text{ in } [0, \tau]$$

Estimation requires dealing with the drivers of censoring:

- Approach 1.** Include predictors of missingness among the conditioning factors
- Approach 2.** Estimate inverse probability of censoring weights (IPCW) and combine them with IPTW.

The Simulation Learner

Inspiration: the Rotterdam study [Royston & Lambert, 2009; Sjolander, 2016]

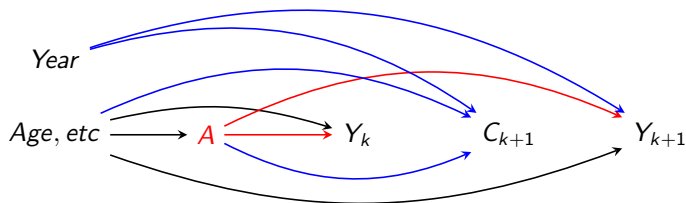
- ▶ - About 3000 women who had undergone surgery for breast cancer and, for some, hormonal therapy was offered in 1978 to 1993
- Outcome of interest: overall mortality
- Strong **negative confounding** of the association between therapy and mortality
- **Informative loss** to follow-up driven by age and year of surgery
- Lack of **positivity** for younger women and those treated before 1985

The Simulation Learner

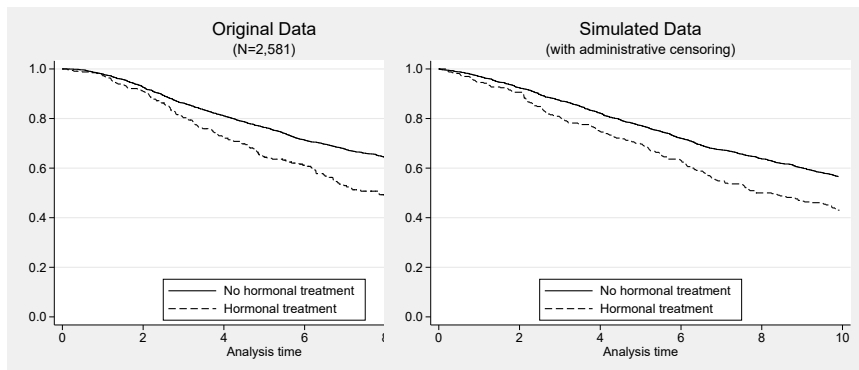
- ▶ Excluded women younger than 40 and with surgery before 1982
- ▶ Retained all the original confounders data

The Simulation Learner

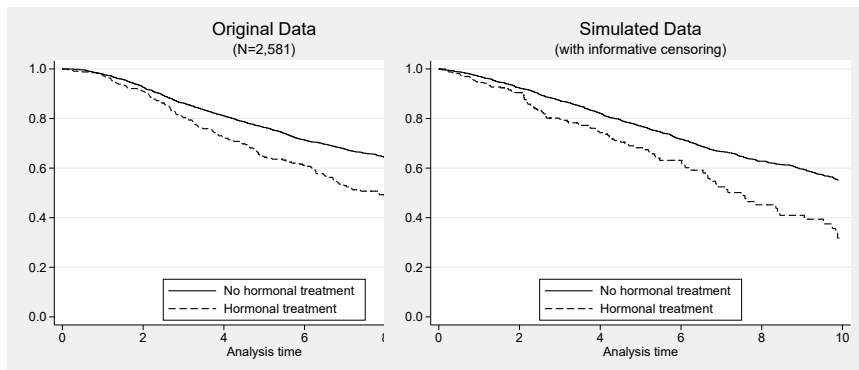
- ▷ Excluded women younger than 40 and with surgery before 1982
- ▷ Retained all the original confounders data
- ▷ Generated:
 - two versions of observed data: **with** / **without** informative censoring
 - potential survival times: $T_{A=0,C=0}$ and $T_{A=1,C=0}$



The Data

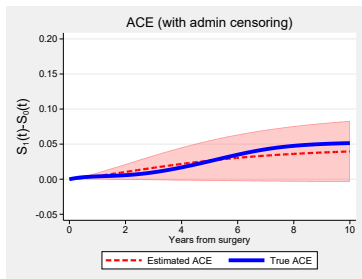


The Data



Results: administrative censoring

1- Model-based marginal counterfactual survival curves

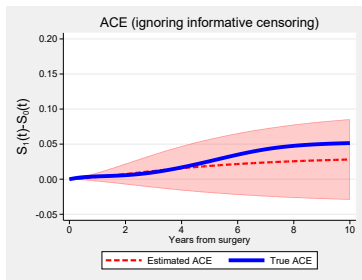


[Estimated in Stata using P. Lambert's `standsurv`]

Time	True ACE(t)	Estimated
5 yr	0.025	0.026 (-0.002, 0.055)
10 yr	0.052	0.040 (-0.004, 0.083)

Results: informative censoring

1- Model-based marginal counterfactual survival curves

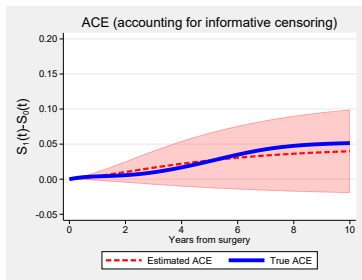


[Estimated in Stata using P. Lambert's `standsurv`]

Time	True ACE(t)	Estimated
5 yr	0.025	0.019 (-0.019, 0.057)
10 yr	0.052	0.028 (-0.029, 0.086)

Results: informative censoring

1- Model-based marginal counterfactual survival curves

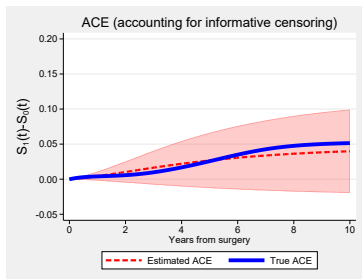


[Estimated in Stata using P. Lambert's `standsurv`]

Time	True ACE(t)	Estimated
5 yr	0.025	0.027
		(-0.013, 0.066)
10 yr	0.052	0.040
		(-0.020, 0.099)

Results: informative censoring

1- Model-based marginal counterfactual survival curves

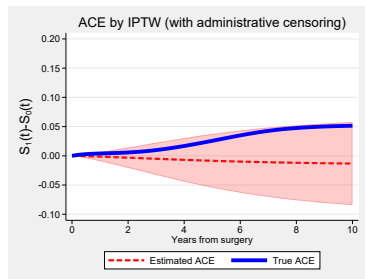


[Estimated in Stata using P. Lambert's `standsurv`]

Time	True RMST	Estimated
5 yr	0.049	0.066 (-0.031, 0.164)
10 yr	0.264	0.240 (-0.114, 0.592)

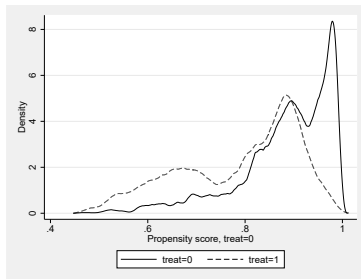
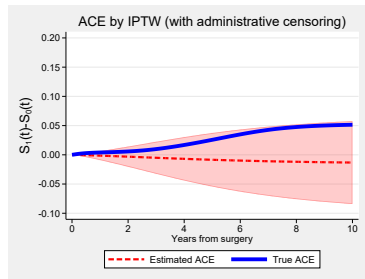
Results: administrative censoring

2- IPTW of MSMs



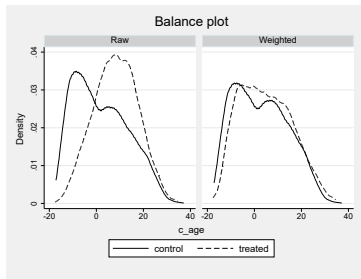
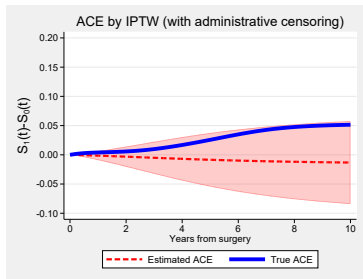
Results: administrative censoring

2- IPTW of MSMs



Results: administrative censoring

2- IPTW of MSMs



Summary

- ▶ Counterfactual-based causal inference has **shifted the focus** from model-based parameters to estimands defined irrespectively of any model.

Summary

- ▶ Counterfactual-based causal inference has **shifted the focus** from model-based parameters to estimands defined irrespectively of any model.
- ▶ This should free us from necessarily wanting to express causal effects on the hazard scale and choose meaningful, clinically relevant quantities.

Summary

- ▶ Counterfactual-based causal inference has **shifted the focus** from model-based parameters to estimands defined irrespectively of any model.
- ▶ This should free us from necessarily wanting to express causal effects on the hazard scale and choose meaningful, clinically relevant quantities.
- ▶ Dealing with censoring calls upon a careful choice of potential outcomes and of appropriate estimation approach.

Summary

- ▶ Counterfactual-based causal inference has **shifted the focus** from model-based parameters to estimands defined irrespectively of any model.
- ▶ This should free us from necessarily wanting to express causal effects on the hazard scale and choose meaningful, clinically relevant quantities.
- ▶ Dealing with censoring calls upon a careful choice of potential outcomes and of appropriate estimation approach.
- ▶ This was just a basic introduction:
the real fun will start with the next speakers!

Summary

- ▶ Counterfactual-based causal inference has **shifted the focus** from model-based parameters to estimands defined irrespectively of any model.
- ▶ This should free us from necessarily wanting to express causal effects on the hazard scale and choose meaningful, clinically relevant quantities.
- ▶ Dealing with censoring calls upon a careful choice of potential outcomes and of appropriate estimation approach.
- ▶ This was just a basic introduction:
the real fun will start with the next speakers!

THANK YOU!

References

- Andersen PK *et al.* (2021) Analysis of time-to-event for observational studies: Guidance to the use of intensity models, *Statistics in Medicine*, 40, 185–211
- Goetghebeur *et al.* (2020) Formulating causal questions and principled statistical answers, *Statistics in Medicine*, 39, 4922–4948
- Paul Lambert's blog:
<https://pclambert.net/software/standsurv/>
- Young *et al.* (2020) A causal framework for classical statistical estimands in failure-time settings with competing events. *Statistics in Medicine* 39: 1199–1236